

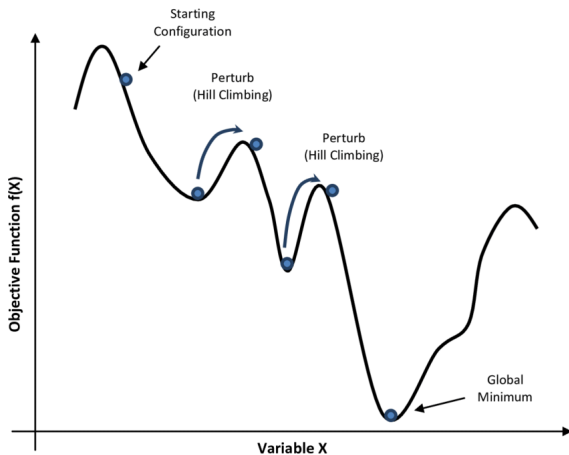
# A Simulated Annealing Algorithm with Constant Temperature for Discrete Stochastic Optimization

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- Original Simulated Annealing Algorithm
- A Modified Simulated Annealing Algorithm for Noisy Functions
- Another Variant of the Modified Simulated Annealing Algorithm

# Original Simulated Annealing Algorithm



# Original Simulated Annealing Algorithm

## Algorithm

### PROCEDURE SIMULATED ANNEALING

begin

**INITIALIZE;**

$M := 0;$

repeat

repeat

**PERTURB**(config.  $i \rightarrow$  config.  $j, \Delta C_{ij}$ );

if  $\Delta C_{ij} \leq 0$  then accept else

if  $\exp(-\Delta C_{ij}/c) > \text{random}[0,1)$  then accept;

if accept then **UPDATE**(configuration  $j$ );

until **equilibrium is approached sufficiently closely;**

$c_{M+1} := f(c_M);$

$M := M + 1;$

until **stop criterion = true (system is 'frozen');**

end.

# Original Simulated Annealing Algorithm

- Most studies have focused on determining an appropriate annealing schedule
- $T_k = \frac{C}{\ln(1+k)}, \forall k \in N$
- The convergence depends on  $C$ , but  $C$  is generally unknown in practice
- Assuming that the objective function values can be evaluated exactly.
- Using the state that is visited by the algorithm in iteration  $k$  as the estimated optimal solution in that iteration

# Modified Simulated Annealing Algorithm

- estimates of the objective function values will be used throughout since exact objective function values are not available
- the state that is visited most often by the algorithm (divided by a normalizer)
- the state that has the best average estimated objective function value
- not require the Markov chain generated by our algorithm to converge to the set of global optimal solutions
- not restrict the temperature to decrease to 0 (more aggressive)

The discrete stochastic optimization problem can be presented as

$$\min_{x \in \varphi} f(x) = \min_{x \in \varphi} E[h(x, Y_x)]$$

where  $Y_x$  is a r.v. Denote  $\varphi^*$  as the set of global optimal solutions.

# The 1st Approach

## Assumptions

**DEFINITION 1.** For each  $x \in \mathcal{S}$ , there exists a subset  $N(x)$  of  $\mathcal{S} \setminus \{x\}$ , which is called the set of neighbors of  $x$ .

**ASSUMPTION 1.** For any  $x, x' \in \mathcal{S}$ ,  $x'$  is reachable from  $x$ ; i.e., there exists a finite sequence  $\{n_i\}_{i=0}^l$  for some  $l$ , such that  $x_{n_0} = x$ ,  $x_{n_l} = x'$ , and  $x_{n_{i+1}} \in N(x_{n_i})$ ,  $i = 0, 1, 2, \dots, l - 1$ .

Let  $R': \mathcal{S} \times \mathcal{S} \rightarrow [0, \infty)$  be a function and for all  $x \in \mathcal{S}$ , define  $D(x) = \sum_{x' \in \mathcal{S}} R'(x, x')$  and

$$R(x, x') = \frac{R'(x, x')}{D(x)}, \quad (4)$$



# The 1st Approach

## Assumptions

**ASSUMPTION 2.** *Let the transition probability  $R(x, x')$  be defined as in Equation (4) and let  $N$  satisfy Definition 1. Then we assume that:*

1.  $R': \mathcal{S} \times \mathcal{S} \rightarrow [0, \infty)$  satisfies  $R'(x, x') > 0 \Leftrightarrow x' \in N(x)$ , and
2.  $R'(x, x') = R'(x', x)$ , for each  $x, x' \in \mathcal{S}$ .

**ASSUMPTION 3.** *The feasible region  $\mathcal{S}$  is a finite set containing at least two states and the set of all global optimal solutions  $\mathcal{S}^*$  is a proper subset of  $\mathcal{S}$ .*

# The 1st Approach

## Assumptions

ASSUMPTION 4. *The temperature  $T$  is a positive (constant) real number.*

ASSUMPTION 5. *Let  $\{L_k\}$  be a sequence of positive integers such that  $L_k \rightarrow \infty$  as  $k \rightarrow \infty$ .*

LEMMA 1. (ANDRADÓTTIR (1995), LEMMA 3.1). *Suppose that  $\{D_k\}$  is a nonhomogeneous Markov chain with a finite state space  $\mathcal{S}$  and with  $P\{D_{k+1} = d \mid D_0, \dots, D_k\} = P_k(D_k, d)$  for all  $d \in \mathcal{S}$  and  $k \in \mathbb{N}$ , where  $P_k(x, x') \rightarrow P(x, x')$  as  $k \rightarrow \infty$  for all  $x, x' \in \mathcal{S}$  and  $P$  is an irreducible and aperiodic Markov matrix. If  $g: \mathcal{S} \rightarrow \mathbb{R}$ , then*

$$(1/K)\sum_{k=1}^K g(D_k) \rightarrow \sum_{d \in \mathcal{S}} \pi_d g(d)$$

*a.s. as  $K \rightarrow \infty$ , where  $\pi$  is the steady-state distribution corresponding to  $P$ .*

# The 1st Approach

Notations:

- $X_k$ : the current state after  $k$  iterations
- $V_k(x)$ : the number of times the Markov chain  $\{X_k\}$  has visited state  $x$  in the first  $k$  iterations for all  $x \in \varphi$
- $X_k^*$ : the state that maximizes  $\frac{V_k(x)}{D(x)}$ , where  $D(x)$  is the normalizer

# The 1st Approach

## Algorithm

**Parameters:**  $R, N, T, \{L_k\}$ .

**Step 0:** Select a starting point  $X_0 \in \mathcal{S}$ . Let  $V_0(X_0) = 1$  and  $V_0(x) = 0$ , for all  $x \in \mathcal{S}, x \neq X_0$ . Let  $k = 0$  and  $X_k^* = X_0$ .

**Step 1:** Given  $X_k = x$ , choose  $Z_k \in N(x)$  such that  $P[Z_k = z \mid X_k = x] = R(x, z)$  for all  $z \in N(x)$ , where  $N(\cdot)$  is defined in Definition 1 and  $R(\cdot, \cdot)$  is given in Equation (4).

**Step 2:** Given  $X_k = x$  and  $Z_k = z$ , generate independent, identically distributed, and unbiased observations  $Y_z(1), Y_z(2), \dots, Y_z(L_k)$  of  $Y_z$  and  $Y_x(1), Y_x(2), \dots, Y_x(L_k)$  of  $Y_x$ . Compute  $\hat{f}_{L_k}(x)$  and  $\hat{f}_{L_k}(z)$  as follows:

$$\hat{f}_{L_k}(s) = \frac{1}{L_k} \sum_{i=1}^{L_k} h(s, Y_s(i)) \quad \text{for } s = x, z. \quad (5)$$

# The 1st Approach

## Algorithm

**Step 3:** Given  $X_k = x$  and  $Z_k = z$ , generate  $U_k \sim U[0, 1]$ , and set

$$X_{k+1} = \begin{cases} Z_k & \text{if } U_k \leq G_{x,z}(k), \\ X_k & \text{otherwise,} \end{cases}$$

where

$$G_{x,z}(k) = \exp\left[\frac{-[\hat{f}_{L_k}(z) - \hat{f}_{L_k}(x)]^+}{T}\right]. \quad (6)$$

**Step 4:** Let  $k = k + 1$ ,  $V_k(X_k) = V_{k-1}(X_k) + 1$ , and  $V_k(x) = V_{k-1}(x)$ , for all  $x \in \mathcal{S}$ ,  $x \neq X_k$ . If  $V_k(X_k)/D(X_k) > V_k(X_{k-1}^*)/D(X_{k-1}^*)$ , then let  $X_k^* = X_k$ ; otherwise let  $X_k^* = X_{k-1}^*$ . Go to Step 1.

# The 1st Approach

## Transition Probability Matrices

$$P_k(x, x') = P\{X_{k+1} = x' \mid X_k = x\}$$
$$= \begin{cases} R(x, x') P\{U_k \leq G_{x, x'}(k)\} & \text{if } x' \in N(x), \\ 1 - \sum_{y \in N(x)} P_k(x, y) & \text{if } x' = x, \\ 0 & \text{otherwise} \end{cases}$$

# The 1st Approach

## Transition Probability Matrices And Stationary Distribution

$$P(x, x')$$

$$= \begin{cases} R(x, x') \exp\left[\frac{-[f(x') - f(x)]^+}{T}\right] & \text{if } x' \in N(x), \\ 1 - \sum_{y \in N(x)} P(x, y) & \text{if } x' = x, \\ 0 & \text{otherwise,} \end{cases} \quad (8)$$

and let  $\pi_x$  be defined as follows:

$$\pi_x = \frac{D(x) \exp\left[\frac{-f(x)}{T}\right]}{\sum_{x' \in \mathcal{G}} D(x') \exp\left[\frac{-f(x')}{T}\right]} \quad (9)$$



# The 1st Approach

## Proposition 2

### Proposition 2

Under Assumptions 1 through 3, the transition probability matrix  $P$  given in Equation (8) is irreducible and aperiodic and has stationary distribution  $\pi$ , where  $\pi$  is a vector whose entries  $\pi_x$  are given in Equation (9).

# The 1st Approach

## Proposition 2

**PROOF.** The proof of irreducibility follows directly from Assumptions 1 and 2 and Equation (8). The proof that  $\pi$  is the stationary distribution of  $P$  can be found in Proposition 3.1 in Mitra et al. (1986). To prove aperiodicity, note that by Assumption 3,  $\mathcal{S}^* \neq \mathcal{S}$ , so by Assumption 1 there exist  $x^* \in \mathcal{S}^*$  and  $x \in N(x^*)$  with  $f(x^*) < f(x)$ . Therefore, from the definition of  $P$  given in Equation (8) and the definition of  $R$  given in Equation (4),  $P(x^*, x^*) > 0$  and therefore,  $P$  is aperiodic.  $\square$

# The 1st Approach

## Proposition 3

### Proposition 3

Suppose that Assumptions 4 and 5 are satisfied and that  $\varphi$  is finite. Then  $P_k(x, x') \rightarrow P(x, x')$  as  $k \rightarrow \infty$  for all  $x, x' \in \varphi$ , where the Markov matrices  $P_k$  and  $P$  are given in Equations (7) and (8), respectively

# The 1st Approach

## Proposition 3

Proof:

$$\begin{aligned} & \lim_{k \rightarrow \infty} P_k(x, x') \\ &= R(x, x') \lim_{k \rightarrow \infty} E \left[ \exp \left[ \frac{-[\hat{f}_k(x') - \hat{f}_k(x)]^+}{T} \right] \right] \\ &= R(x, x') E \left[ \lim_{k \rightarrow \infty} \exp \left[ \frac{-[\hat{f}_k(x') - \hat{f}_k(x)]^+}{T} \right] \right] \\ &= R(x, x') E \left[ \exp \left[ \frac{-[f(x') - f(x)]^+}{T} \right] \right] = P(x, x') \end{aligned}$$

# The 1st Approach

## Theorem 4

### Theorem 4

Under Assumptions 1 through 5 , the sequence  $\{X_k^*\}$  generated by Algorithm 1 converges almost surely to the set  $\varphi^*$  (in the sense that there exists a set  $A$  such that  $P(A) = 1$  and for all  $\omega \in A$ , there exists  $K_\omega > 0$  such that  $X_k^*(\omega) \in \varphi^*$  for all  $k \geq K_\omega$  ).

# The 1st Approach

## Theorem 4

Proof: By Lemma 1 (let  $D_k = X_k$  for all  $k \in \mathbb{N}$  and  $g(d) = I_{\{d=x\}}$  for all  $d \in \varphi$ ), we have that

$$\frac{V_k(x)}{kD(x)} = \frac{1}{D(x)} \times \frac{1}{k} \sum_{i=0}^k I_{\{X_i=x\}} \rightarrow \frac{\pi_x}{D(x)} \quad \text{a.s. as } k \rightarrow \infty$$

for all  $x \in \varphi$ , where  $V_k(x), x \in \varphi$ . It is clear that  $\pi_x > 0$  for all  $x \in \varphi$  so that Equation (9) yields

$$\frac{\pi_y/D(y)}{\pi_x/D(x)} = \exp \left[ \frac{-[f(y) - f(x)]}{T} \right]$$

for all  $x, y \in \varphi$ . Therefore,  $\pi_y/D(y) \leq \pi_x/D(x)$  if and only if  $f(y) \geq f(x)$ . This shows that

$$\arg \max_{x \in \varphi} \frac{\pi_x}{D(x)} = \varphi^*$$

# The 2nd Approach

- This variant uses the same mechanism for moving around the state space as Algorithm 1 but a different approach for estimating the optimal solution.
- Selecting the state with the best average estimated objective function value obtained from all the previous estimates to be the estimated optimal solution

Assumption 5':  $\{L_k\}$  is a sequence of positive integers satisfying  
 $\lim_{k \rightarrow \infty} L_k = L \leq \infty$

# The 2nd Approach

## Algorithm

**Parameters:**  $R, N, T, \{L_k\}$ .

**Step 0:** Select a starting point  $X_0 \in \mathcal{S}$ . For all  $x \in \mathcal{S}$ , let  $A_0(x) = 0$  and  $C_0(x) = 0$ . Let  $k = 0$  and  $X_k^* = X_0$ .

**Step 1:** Identical to Step 1 of Algorithm 1.

**Step 2:** Given  $X_k = x$  and  $Z_k = z$ , generate independent, identically distributed, and unbiased observations  $Y_z(1), Y_z(2), \dots, Y_z(L_k)$  of  $Y_z$  and  $Y_x(1), Y_x(2), \dots, Y_x(L_k)$  of  $Y_x$ . Compute  $\hat{f}_{L_k}(x)$  and  $\hat{f}_{L_k}(z)$  as in Equation (5) for  $s = x, z$ . Let  $C_{k+1}(s) = C_k(s) + L_k$  and  $A_{k+1}(s) = A_k(s) + L_k \hat{f}_{L_k}(s)$  for  $s = x, z$ . Moreover, let  $C_{k+1}(x') = C_k(x')$  and  $A_{k+1}(x') = A_k(x')$  for all  $x' \in \mathcal{S}, x' \neq x, z$ .

**Step 3:** Identical to Step 3 of Algorithm 1.

**Step 4:** Let  $k = k + 1$  and select  $X_k^* \in \arg \min_{x \in \mathcal{S}} A_k(x)/C_k(x)$  (use the convention  $0/0 = +\infty$ ). Go to Step 1.



# The 2nd Approach

## Theorem 5

### Theorem 5

Suppose that Assumptions 1, 2, 3, 4 and 5' are satisfied, the sequence  $\{X_k^*\}$  generated by Algorithm 2 converges almost surely to the set  $\varphi^*$  (in the sense that there exists a set  $A$  such that  $P(A) = 1$  and for all  $\omega \in A$ , there exists  $K_\omega > 0$  such that  $X_k^*(\omega) \in \varphi^*$  for all  $k \geq K_\omega$ ).

# The 2nd Approach

## Theorem 5

Proof:

(i)  $L = \infty$

$P_k \rightarrow P$ ,  $P$  is irreducible and aperiodic and has the stationary distribution  $\pi$ . Since  $\varphi$  is finite,  $\pi_x > 0$  for all  $x \in \varphi$ .

By Lemma 1,  $\frac{V_k(x)}{k} \rightarrow \pi_x > 0$  a.s. as  $k \rightarrow \infty$  for for all  $x \in \varphi$

$\Rightarrow V_k(x) \rightarrow \infty$  as  $k \rightarrow \infty$  for for all  $x \in \varphi$ . Clearly,  $C_k(x) \geq V_k(x) - 1$

$\Rightarrow C_k(x) \rightarrow \infty$  a.s. as  $k \rightarrow \infty$  for for all  $x \in \varphi$

$\Rightarrow \frac{A_k(x)}{C_k(x)} \rightarrow f(x)$  a.s. as  $k \rightarrow \infty$  for for all  $x \in \varphi$

# The 2nd Approach

## Theorem 5

Proof:

(ii)  $L < \infty$ ,  $P_k(x, x') \rightarrow P'(x, x')$

Firstly, we show that  $P'$  is irreducible

$$\begin{aligned} P'(x, x') &= P\{X_{k+1} = x' \mid X_k = x\} \\ &= \begin{cases} R(x, x')p'_{x,x'} & x' \in N(x), \\ 1 - \sum_{y \in N(x)} P'(x, y) & \text{if } x' = x, \\ 0 & \text{otherwise,} \end{cases} \quad (11) \end{aligned}$$

$$\begin{aligned} p'_{x,x'} &= E[\exp[-[\hat{f}(x') - \hat{f}(x)]^+ / T]], \text{ and } \hat{f}(s) \\ &= (1/L) \sum_{i=1}^L h(s, Y_s(i)) \text{ for } s = x, x'. \end{aligned}$$

By Jensen's inequality, we have that

$$\begin{aligned} p'_{x,x'} &\geq \exp\left[\frac{-1}{T} E[\hat{f}(x') - \hat{f}(x)]^+\right] \\ &\geq \exp\left[\frac{-1}{T} E[|\hat{f}(x')| + |\hat{f}(x)|]\right] \\ &> 0, \end{aligned}$$

Then, we show that  $P'$  is aperiodic

$$P'(x, x) = 0, \forall x \in \mathcal{S}$$

$$\Leftrightarrow \sum_{x' \in N(x)} P'(x, x') = 1, \forall x \in \mathcal{S}$$

$$\Leftrightarrow E \left[ \exp \left[ \frac{-[\hat{f}(x') - \hat{f}(x)]^+}{T} \right] \right] = 1,$$

$$\forall x \in \mathcal{S}, x' \in N(x)$$

$$\Leftrightarrow \exp \left[ \frac{-[\hat{f}(x') - \hat{f}(x)]^+}{T} \right] = 1 \text{ a.s.},$$

$$\forall x \in \mathcal{S}, x' \in N(x)$$

$$\Leftrightarrow \hat{f}(x') - \hat{f}(x) \leq 0 \text{ a.s.}, \forall x \in \mathcal{S}, x' \in N(x)$$

$$\Leftrightarrow \hat{f}(x') = \hat{f}(x) \text{ a.s.}, \forall x \in \mathcal{S}, x' \in N(x)$$

$$\Leftrightarrow \hat{f}(x') = \hat{f}(x) \text{ a.s.}, \forall x, x' \in \mathcal{S}$$

$$\Leftrightarrow f(x') = f(x), \forall x, x' \in \mathcal{S}.$$

Since this contradicts Assumption 3, we have shown that  $P'$  is aperiodic.

Since  $P_k \rightarrow P'$  as  $k \rightarrow \infty$ , where  $P'$  is irreducible and aperiodic and  $P'$  has a stationary distribution  $r$ , with  $r_x > 0$  for all  $x \in \mathcal{S}$ , by Lemma 1,  $V_k(x)/k \rightarrow r_x > 0$  a.s. as  $k \rightarrow \infty$  for all  $x \in \mathcal{S}$ . This implies that  $V_k(x) \rightarrow \infty$  a.s. as  $k \rightarrow \infty$  for all  $x \in \mathcal{S}$ . The rest of the proof is similar to the proof for the case  $L = \infty$ .  $\square$

- Algorithm 2 will converge to the set of global optimal solutions to the optimization problem more rapidly than Algorithm 1
- Algorithm 2 is expected to obtain a good estimate of the optimal solution quickly. Since Algorithm 2 selects the state in  $\varphi$  that has the best objective function value among the states that have been visited by the algorithm so far to be the estimate of the optimal solution.

Thanks!