Bayesian Optimization for Simulation Optimization of Continuous Parameters

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Outline



- 2 Bayesian Optimization
- **3** Acquisition Functions for Sampling
- 4 Remaining Issues



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Outline

1 Introduction

- 2 Bayesian Optimization
- 3 Acquisition Functions for Sampling

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- 4 Remaining Issues
- 5 Summary

Simulation Optimization



- Evaluate the performance of a system design.
- 2 Select the best design $\mathbf{x}^* = \arg \min_{\mathbf{x}} y(\mathbf{x})$.

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Types of Design Variables and Samples



- Sample types:
 - Noiseless
 - 2 Noisy
- Variable types.
- Today's topic: Continuous variables + noiseless/noisy samples

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Study Goal: Bayesian optimization for simulation optimization of continuous parameters.

- Noiseless samples:
 - Jones D R, Schonlau M, Welch W J. Efficient global optimization of expensive black-box functions[J]. *Journal of Global optimization*, 1998, 13(4): 455-492. (citations: 6538)
- 2 Noisy samples:
 - Scott W, Frazier P, Powell W. The correlated knowledge gradient for simulation optimization of continuous parameters using gaussian process regression[J]. *SIAM Journal on Optimization*, 2011, 21(3): 996-1026.

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Components of Bayesian Optimization

- A typical Bayesian optimization consists of two parts:
 - Gaussian process (stochastic kriging): predicting function values.
 - Sampling methods: determining the design point that should be sampled. (This slides discuss two acquisition functions (figure of merit): EI and KG.)

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Outline

Introduction

Bayesian Optimization

- Gaussian process
- Bayesian Learning
- Sampling Methods
- 3 Acquisition Functions for Sampling
 - Simulation Optimization with Noiseless Samples
 - Simulation Optimization with Noisy Samples
 - Calculation
- Remaining Issues
 - Model Validation
 - Identifying Important Factors

Summary

Surrogates

- Goal: to predict surface values $y(\mathbf{x}), \mathbf{x} \in \mathcal{X}$, given a limited number of random observations $Y(\mathbf{x}^i), i = 0, 1, \ldots, n$.
- Typical surrogates:
 - Linear basis function models.
 - ② Gaussian process.
- Relationship: they can be unified through the ridge regularization (Hong and Zhang, 2021).

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Two examples



- GP regression can capture fluctuated surfaces easily.
- However, linear basis function model can also achieve it.
- Advantage of GP: Bayesian learning.

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An example (Continued)



• Bayesian learning: construct Bayesian credible region (analogous to confidence intervals in frequentist statistics).

Outline

Introduction

Bayesian Optimization

- Gaussian process
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- Sampling Methods
- 3 Acquisition Functions for Sampling
 - Simulation Optimization with Noiseless Samples
 - Simulation Optimization with Noisy Samples
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Summary

Learning Targets



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Subjective v.s. Objective. Let's flip a coin:

- Objective: the probability that a coin has a head is 1/2 (or 1/3 if the coin is uneven).
- Subjective: I think the probability that a coin has a head is within [1/3, 2/3] and has an uniform distribution.

Bayesian prior and posterior are subjective probabilities.

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• Suppose we want to learn the value of $y(\mathbf{x}^0)$ (for short, $y(\mathbf{x}^0) := y$). The random sample satisfies $E[Y_i] = y$.



- The unknown value y is regarded as a random variable.
- Subjective belief: $P(y \le y_1) = 0.2$, $P(y \le y_2) = 0.5$, $P(y \le y_3) = 0.7$, Can be a distribution.
- We sequentially update the belief toward its value: $\Pi^0, \Pi^1, \ldots, \Pi^T$.

• Filtration:



• Example: Given y, the sample Y_1 follows $\mathcal{N}(y, 1/\tau)$ (τ is known).

- Before observing Y_1 , note that $Y_1 = y + (Y_1 y)$.
- In our belief, the prior of y as $\mathcal{N}(\mu^0, 1/\tau^0)$.
- Sample noise $Y_1 y \sim \mathcal{N}(0, 1/\tau)$, independent of y's belief.
- So, in our belief, $Y_1 \sim \mathcal{N}(\mu^0, 1/\tau^0 + 1/\tau)$. (predictive distribution of Y_1)

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Image: A matrix

- Example: Given y, the sample follows $\mathcal{N}(y, 1/\tau)$ and τ is known.
 - Given \mathcal{F}^0 and before observing Y_1 , the joint distribution of $(y, Y_1)^\top$:

$$(y, Y_1)^{\top} \sim \mathcal{N}\left((\mu^0, \mu^0)^{\top}, \begin{pmatrix} 1/\tau^0 & 1/\tau^0 \\ 1/\tau^0 & 1/\tau^0 + 1/\tau \end{pmatrix}\right).$$

Hint:

 $\operatorname{Cov}(y, Y_1) = \operatorname{Cov}(y, y + (Y_1 - y)) = \operatorname{Cov}(y, y) + \operatorname{Cov}(y, Y_1 - y) = 1/\tau^0.$ • The conditional distribution of y given Y_1 is

$$y|Y_1 \sim \mathcal{N}\left(\frac{\tau^0\mu^0 + \tau Y_1}{\tau^0 + \tau}, \frac{1}{\tau + \tau^0}\right) \triangleq \mathcal{N}(\mu^1, 1/\tau^1).$$

(the posterior distribution of y after observing Y_1)

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• Example: Given y, the sample follows $\mathcal{N}(y, 1/\tau)$ and τ is known.

- Given $\Pi^0 = \mathcal{N}(\mu^0, 1/\tau^0), \ \mu^1 = \frac{\tau^0 \mu^0 + \tau Y_1}{\tau^0 + \tau}$ is random due to Y_1 .
- Note that the predictive distribution of Y_1 is

$$\mathcal{N}(\mu^0, 1/\tau^0 + 1/\tau).$$

• The predictive distribution of $\mu^1 = \frac{\tau^0 \mu^0}{\tau^0 + \tau} + \frac{\tau}{\tau^0 + \tau} Y_1$ given \mathcal{F}^0 is

$$\mathcal{N}\left(\mu^{0}, \frac{\tau}{\tau^{0}(\tau^{0}+\tau)}
ight).$$

- In summary, we have learned given \mathcal{F}^0 ,
 - the predictive distribution of Y_1 and μ^1 before observing Y_1 ,
 - the posterior distribution of y ($\Pi^1 = \mathcal{N}(\mu^1, 1/\tau^1)$) after observing Y_1 .

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- Example: Given y, the sample follows $\mathcal{N}(y, 1/\tau)$ and τ is known.
 - Given \mathcal{F}^1 and before observing Y_2 ,
 - the predictive distribution of Y_2 is $\mathcal{N}(\mu^1, 1/\tau^1 + 1/\tau)$.
 - the predictive distribution of μ^2 is $\mathcal{N}\left(\mu^1, \frac{\tau}{\tau^1(\tau^1+\tau)}\right)$
 - Given \mathcal{F}^2 (given \mathcal{F}^1 and after observing Y_2), the posterior distribution of y is $\mathcal{N}(\mu^2, 1/\tau^2)$, where

$$\begin{aligned} \tau^2 = &\tau + \tau^1 = 2\tau + \tau^0, \\ \mu^2 = &\frac{\tau^1 \mu^1 + \tau Y_2}{\tau^1 + \tau} = \frac{\tau^1 \frac{\tau^0 \mu^0 + \tau Y_1}{\tau^0 + \tau} + \tau Y_2}{2\tau + \tau^0} = \frac{\tau^0 \mu^0 + \tau Y_1 + \tau Y_2}{2\tau + \tau^0}. \end{aligned}$$

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• Example: Given y, the sample follows $\mathcal{N}(y, 1/\tau)$ and τ is known.

- Given \mathcal{F}^{n-1} and before observing Y_n ,
 - the predictive distribution of Y_n is $\mathcal{N}(\mu^{n-1}, 1/\tau^{n-1} + 1/\tau)$.
 - the predictive distribution of μ^n is $\mathcal{N}\left(\mu^{n-1}, \frac{\tau}{\tau^{n-1}(\tau^{n-1}+\tau)}\right)$
- Given \mathcal{F}^n (given \mathcal{F}^{n-1} and after observing Y_n), the posterior distribution of y is $\mathcal{N}(\mu^n, 1/\tau^n)$, where

$$\tau^{n} = \tau + \tau^{n-1} = n\tau + \tau^{0},$$

$$\mu^{n} = \frac{\tau^{n-1}\mu^{n-1} + \tau Y_{1}}{\tau^{n-1} + \tau} = \frac{\tau^{0}\mu^{0} + \tau (Y_{1} + Y_{2} + \dots + Y_{n})}{n\tau + \tau^{0}}.$$

• Question 1: what if the observation has no noise, i.e., $\tau = \infty$? $(y|y \sim \mathcal{N}(y,0))$

Conjugate Family II

- Conjugacy: the posterior distribution is in the same family as the prior distribution.
- Other conjugate families:

Sampling Distribution	Conjugate Family
Exponential	Gamma
Poisson	Gamma
Uniform	Pareto
Bernoulli	Beta
Normal with unknown variance	Normal-Gamma

- Non-conjugate priors: posterior may not be analytically tractable.
 - numerical calculation: Markov chain Monte Carlo, importance sampling.

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Bayesian Learning of Multivariate Normal

• A vector $\mathbf{y} = (y(\mathbf{x}^1), y(\mathbf{x}^2), \dots, y(\mathbf{x}^d))^\top$ to estimate.



- The sample at \mathbf{x}^i is normal: $\mathcal{N}(y(\mathbf{x}^i), \lambda(\mathbf{x}^i))$.
- Set the prior Π^0 as $\mathcal{N}(\mu^0, \Sigma^0)$. The posterior distribution Π^1 after observing $Y(\mathbf{x}^{i_0})$ is $\mathcal{N}(\mu^1, \Sigma^1)$ where

$$\boldsymbol{\mu}^{1} = \boldsymbol{\mu}^{0} + \Sigma^{0} e_{\mathbf{x}^{i_{0}}} \left(\Sigma^{0}(\mathbf{x}^{i_{0}}, \mathbf{x}^{i_{0}}) + \lambda(\mathbf{x}^{i_{0}}) \right)^{-1} \left(Y(\mathbf{x}^{i_{0}}) - \mu^{0}(\mathbf{x}^{i_{0}}) \right)$$

$$\Sigma^{1} = \Sigma^{0} - \Sigma^{0} e_{\mathbf{x}^{i_{0}}} \left(\Sigma^{0}(\mathbf{x}^{i_{0}}, \mathbf{x}^{i_{0}}) + \lambda(\mathbf{x}^{i_{0}}) \right)^{-1} \left(\Sigma^{0} e_{\mathbf{x}^{i_{0}}} \right)^{\top}$$

• Hint: $\mathbf{\Pi}^1 = (y(\mathbf{x}^1), y(\mathbf{x}^2), \dots, y(\mathbf{x}^d))^\top | Y(\mathbf{x}^{i_0}).$

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Answer to Questions for Noiseless Sampling

• Question 2: what if the observation has no noise, i.e., $\lambda(\mathbf{x}^n) = 0$? Prior:

$$(y(\mathbf{x}^0), y(\mathbf{x}))^\top \sim \mathcal{N}\left(\left(\begin{array}{c} \boldsymbol{\mu}^0(\mathbf{x}^0) \\ \boldsymbol{\mu}^0(\mathbf{x}) \end{array} \right), \left(\begin{array}{cc} \boldsymbol{\Sigma}^0(\mathbf{x}^0, \mathbf{x}^0) & \boldsymbol{\Sigma}^0(\mathbf{x}, \mathbf{x}^0) \\ \boldsymbol{\Sigma}^0(\mathbf{x}^0, \mathbf{x}) & \boldsymbol{\Sigma}^0(\mathbf{x}, \mathbf{x}) \end{array} \right) \right)$$

Posterior:

$$\begin{array}{l} (y(\mathbf{x}^{0}), y(\mathbf{x}))^{\top} | y(\mathbf{x}^{0}) \sim \\ \mathcal{N} \left(\begin{pmatrix} y(\mathbf{x}^{0}) \\ \mu^{0}(\mathbf{x}) + \frac{\Sigma^{0}(\mathbf{x}, \mathbf{x}^{0})}{\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}^{0})} \left(y(\mathbf{x}^{0}) - \mu^{0}(\mathbf{x}^{0}) \right) \end{array} \right), \begin{pmatrix} 0 & 0 \\ 0 & \Sigma^{0}(\mathbf{x}, \mathbf{x}) - \frac{(\Sigma^{0}(\mathbf{x}, \mathbf{x}^{0}))^{2}}{\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}^{0})} \end{array} \right) \right)$$



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Gaussian Process

- A function $y(\mathbf{x}), \mathbf{x} \in \mathcal{X}$, to estimate.
- A prior on $y(\mathbf{x}), \mathbf{x} \in \mathcal{X}$, is a Gaussian process if the prior of any $(y(\mathbf{x}^0), y(\mathbf{x}^1), \dots, y(\mathbf{x}^n))^{\top}$ has a multivariate Gaussian distribution.
 - Mean function: $\mu^0(\mathbf{x})$
 - Covariance function: $\Sigma^0(\mathbf{x}, \mathbf{x}') = Cov(\mu^0(\mathbf{x}), \mu^0(\mathbf{x}')).$
 - The prior on $(y(\mathbf{x}^0), y(\mathbf{x}^1), \dots, y(\mathbf{x}^n))^{\dagger}$ is multivariate Gaussian
 - mean: $\mu^0([\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^n])$: $(\mu^0(\mathbf{x}^0), \mu^0(\mathbf{x}^1), \dots, \mu^0(\mathbf{x}^n))^\top$ covariance matrix $\Sigma^0([\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^n])$:

$$\left(\begin{array}{ccccc} \Sigma^{0}(\mathbf{x}^{0},\mathbf{x}^{0}) & \Sigma^{0}(\mathbf{x}^{0},\mathbf{x}^{1}) & \dots & \Sigma^{0}(\mathbf{x}^{0},\mathbf{x}^{n}) \\ \Sigma^{0}(\mathbf{x}^{1},\mathbf{x}^{0}) & \Sigma^{0}(\mathbf{x}^{1},\mathbf{x}^{1}) & \dots & \Sigma^{0}(\mathbf{x}^{1},\mathbf{x}^{n}) \\ \vdots & \vdots & \ddots & \vdots \\ \Sigma^{0}(\mathbf{x}^{n},\mathbf{x}^{0}) & \Sigma^{0}(\mathbf{x}^{n},\mathbf{x}^{1}) & \dots & \Sigma^{0}(\mathbf{x}^{n},\mathbf{x}^{n}) \end{array}\right)$$



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- For any $\mathbf{x} \in \mathcal{X}$, the prior distribution on $(y(\mathbf{x}^0), y(\mathbf{x}^1))^{\top}$ is $\mathcal{N}(\boldsymbol{\mu}^0([\mathbf{x}^0, \mathbf{x}^1]), \boldsymbol{\Sigma}^0([\mathbf{x}^0, \mathbf{x}^1])).$
- Given a sample $\hat{y}^1 = Y(\mathbf{x}^0)$ at \mathbf{x}^0 . The posterior distribution of $(y(\mathbf{x}^0), y(\mathbf{x}^1)))^{\top}$ is $\mathcal{N}(\boldsymbol{\mu}^1([\mathbf{x}^0, \mathbf{x}^1]), \Sigma^1([\mathbf{x}^0, \mathbf{x}^1]))$ where

$$\begin{split} & \boldsymbol{\mu}^{1}([\mathbf{x}^{0},\mathbf{x}^{1}]) \\ = & \boldsymbol{\mu}^{0}([\mathbf{x}^{0},\mathbf{x}^{1}]) + \boldsymbol{\Sigma}^{0}([\mathbf{x}^{0},\mathbf{x}^{1}])e_{\mathbf{x}^{0}}\left(\boldsymbol{\Sigma}^{0}(\mathbf{x}^{0},\mathbf{x}^{0}) + \boldsymbol{\lambda}(\mathbf{x}^{0})\right)^{-1}\left(\hat{y}^{1} - \boldsymbol{\mu}^{0}(\mathbf{x}^{0})\right), \\ & \boldsymbol{\Sigma}^{1}([\mathbf{x}^{0},\mathbf{x}^{1}]) \\ = & \boldsymbol{\Sigma}^{0}([\mathbf{x}^{0},\mathbf{x}^{1}]) - \boldsymbol{\Sigma}^{0}([\mathbf{x}^{0},\mathbf{x}^{1}])e_{\mathbf{x}^{0}}\left(\boldsymbol{\Sigma}^{0}(\mathbf{x}^{0},\mathbf{x}^{0}) + \boldsymbol{\lambda}(\mathbf{x}^{0})\right)^{-1}e_{\mathbf{x}^{0}}^{\top}\boldsymbol{\Sigma}^{0}([\mathbf{x}^{0},\mathbf{x}^{1}]). \end{split}$$

• Noiseless:
$$y(\mathbf{x}^1)|y(\mathbf{x}^0) \sim \mathcal{N}\left(\mu^0(\mathbf{x}^1) + \frac{\Sigma^0(\mathbf{x}^1, \mathbf{x}^0)}{\Sigma^0(\mathbf{x}^0, \mathbf{x}^0)} \left(y(\mathbf{x}^0) - \mu^0(\mathbf{x}^0)\right), \Sigma^0(\mathbf{x}^1, \mathbf{x}^1) - \frac{(\Sigma^0(\mathbf{x}^1, \mathbf{x}^0))^2}{\Sigma^0(\mathbf{x}^0, \mathbf{x}^0)}\right)$$

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• More generally, after collecting $\hat{y}^{n+1} = Y(\mathbf{x}^n)$ (given \mathcal{F}^{n+1}), the posterior mean of $(y(\mathbf{x}^0), y(\mathbf{x}^1), \dots, y(\mathbf{x}^n))^{\top}$ in the recursive form is

$$\begin{pmatrix} \mu^{n+1}(\mathbf{x}^{0}) \\ \mu^{n+1}(\mathbf{x}^{1}) \\ \vdots \\ \mu^{n+1}(\mathbf{x}^{n}) \end{pmatrix} = \begin{pmatrix} \mu^{n}(\mathbf{x}^{0}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{n}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{n}) \end{pmatrix} + \Sigma^{n}([\mathbf{x}^{0}, \dots, \mathbf{x}^{n}])e_{\mathbf{x}^{n}} \left(\Sigma^{n}(\mathbf{x}^{n}, \mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})\right)^{-1} \left(\hat{y}^{n+1} - \mu^{n}(\mathbf{x}^{n})\right) \\ = \begin{pmatrix} \mu^{n}(\mathbf{x}^{0}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{n}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{n}) \end{pmatrix} + \frac{\Sigma^{n}([\mathbf{x}^{0}, \dots, \mathbf{x}^{n}])e_{\mathbf{x}^{n}}}{\sqrt{\Sigma^{n}(\mathbf{x}^{n}, \mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})} \frac{\hat{y}^{n+1} - \mu^{n}(\mathbf{x}^{n})}{\sqrt{\Sigma^{n}(\mathbf{x}^{n}, \mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})} \\ & = \begin{pmatrix} \mu^{n}(\mathbf{x}^{0}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{1}) \\ \vdots \\ \mu^{n}(\mathbf{x}^{n}) \end{pmatrix} + \tilde{\sigma}(\Sigma^{n}([\mathbf{x}^{0}, \dots, \mathbf{x}^{n}]), \mathbf{x}^{n}) \frac{\hat{y}^{n+1} - \mu^{n}(\mathbf{x}^{n})}{\sqrt{\Sigma^{n}(\mathbf{x}^{n}, \mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})}$$

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• Before collecting $\hat{y}^{n+1} = Y(\mathbf{x}^n)$ and given \mathcal{F}^n ,

$$\hat{y}^{n+1} = y(\mathbf{x}^n) + (Y(\mathbf{x}^n) - y(\mathbf{x}^n)) \sim \mathcal{N}(\boldsymbol{\mu}^n(\mathbf{x}^n), \boldsymbol{\Sigma}^n(\mathbf{x}^n, \mathbf{x}^n) + \boldsymbol{\lambda}(\mathbf{x}^n)).$$

• the predictive distribution of $(\boldsymbol{\mu}^{n+1}(\mathbf{x}^0), \boldsymbol{\mu}^{n+1}(\mathbf{x}^1), \dots, \boldsymbol{\mu}^{n+1}(\mathbf{x}^n))^{\top}$

$$\begin{pmatrix} \boldsymbol{\mu}^{n+1}(\mathbf{x}^{0}) \\ \boldsymbol{\mu}^{n+1}(\mathbf{x}^{1}) \\ \vdots \\ \boldsymbol{\mu}^{n+1}(\mathbf{x}^{n}) \end{pmatrix} = \begin{pmatrix} \boldsymbol{\mu}^{n}(\mathbf{x}^{0}) \\ \boldsymbol{\mu}^{n}(\mathbf{x}^{1}) \\ \vdots \\ \boldsymbol{\mu}^{n}(\mathbf{x}^{n}) \end{pmatrix} + \frac{\Sigma^{n}([\mathbf{x}^{0},\dots,\mathbf{x}^{n}])e_{\mathbf{x}^{n}}}{\sqrt{\Sigma^{n}(\mathbf{x}^{n},\mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})}} \frac{\hat{y}^{n+1} - \boldsymbol{\mu}^{n}(\mathbf{x}^{n})}{\sqrt{\Sigma^{n}(\mathbf{x}^{n},\mathbf{x}^{n}) + \lambda(\mathbf{x}^{n})}}$$
$$= \begin{pmatrix} \boldsymbol{\mu}^{n}(\mathbf{x}^{0}) \\ \boldsymbol{\mu}^{n}(\mathbf{x}^{1}) \\ \vdots \\ \boldsymbol{\mu}^{n}(\mathbf{x}^{n}) \end{pmatrix} + \tilde{\sigma}(\Sigma^{n}([\mathbf{x}^{0},\dots,\mathbf{x}^{n}]),\mathbf{x}^{n})Z^{(n+1)} \text{ (Note: } Z^{(n+1)} \sim \mathcal{N}(0,1).)$$

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• After collecting $\hat{y}^{n+1} = Y(\mathbf{x}^n)$ (given \mathcal{F}^{n+1}), the posterior distribution of $y(\mathbf{x})$ in the direct form is $\mathcal{N}(\mu^{n+1}(\mathbf{x}), \Sigma^{n+1}(\mathbf{x}))$ where (hint: $y(\mathbf{x})|\hat{y}^1, \hat{y}^2, \dots, \hat{y}^{n+1}$)

$$\mu^{n+1}(\mathbf{x}) = \mu^{0}(\mathbf{x}) + \left(\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}), \dots, \Sigma^{0}(\mathbf{x}^{n}, \mathbf{x})\right) \left(S^{n}\right)^{-1} \begin{pmatrix} \hat{y}^{1} - \mu^{0}(\mathbf{x}^{0}) \\ \vdots \\ \hat{y}^{n+1} - \mu^{0}(\mathbf{x}^{n}) \end{pmatrix}$$
$$\Sigma^{n+1}(\mathbf{x}) = \Sigma^{0}(\mathbf{x}, \mathbf{x}) - \left(\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}), \dots, \Sigma^{0}(\mathbf{x}^{n}, \mathbf{x})\right) \left(S^{n}\right)^{-1} \begin{pmatrix} \Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}) \\ \vdots \\ \Sigma^{0}(\mathbf{x}^{n}, \mathbf{x}) \end{pmatrix}$$

and
$$S^n = \Sigma^0([\mathbf{x}^0, \dots, \mathbf{x}^n]) + diag([\lambda(\mathbf{x}^0), \dots, \lambda(\mathbf{x}^n)]).$$

• Question 3: what if the observation has no noise?

Answer to Questions for Noiseless Sampling

• Question 3: Note that given $\lambda(\mathbf{x}^i) = 0, i = 0, 1, ..., n$, $S^n = \Sigma^0([\mathbf{x}^0, ..., \mathbf{x}^n])$. Then,

$$(S^n)^{-1} \begin{pmatrix} \Sigma^0(\mathbf{x}^0, \mathbf{x}_i) \\ \vdots \\ \Sigma^0(\mathbf{x}^n, \mathbf{x}_i) \end{pmatrix} = \begin{pmatrix} 0 \\ \vdots \\ 1 \\ \vdots \\ 0 \end{pmatrix} \triangleq e_i$$

• So, $y(\mathbf{x}^i) \sim \mathcal{N}(\mu^{n+1}(\mathbf{x}^i), \Sigma^{n+1}(\mathbf{x}^i))$ where $\mu^{n+1}(\mathbf{x}^i) = y(\mathbf{x}^i)$ and $\Sigma^{n+1}(\mathbf{x}^i) = 0, i = 0, 1, \dots, n.$

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Notations Check Table

This document	KG	EGO
$y(\mathbf{x})$	$\mu(\mathbf{x})$	$y(\mathbf{x})$
$\min_{\mathbf{x}\in\mathcal{X}} y(\mathbf{x})$	$\max_{\mathbf{x}\in\mathcal{X}}\mu(\mathbf{x})$	$\min_{\mathbf{x}\in\mathcal{X}} y(\mathbf{x})$
$\Sigma^0({f x},{f x}')$	$\Sigma^0({f x},{f x}')$	$\sigma^2 \operatorname{Corr}[\epsilon(\mathbf{x}), \epsilon(\mathbf{x}')]$
β	β	σ^2
S^{n-1}	S^{n-1}	$\sigma^2 \mathbf{R}$
$\mu^0(\mathbf{x})$	$\mu^0(\mathbf{x})$	$\sum_{h} \beta_h f_h(\mathbf{x})$ and μ
$\mu^n(\mathbf{x}^*)$	$\mu^n(\mathbf{x}^*)$	$\hat{y}(\mathbf{x}^*)$
\hat{y}^i	\hat{y}^i	$y(\mathbf{x}^{(i+1)})$
$\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^{n-1}$	$\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^{n-1}$	$\mathbf{x}^{(1)}, \mathbf{x}^{(2)} \dots, \mathbf{x}^{(n)}$

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- Gaussian process
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 - Simulation Optimization with Noisy Samples
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 - Identifying Important Factors

Summary

Example 1



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Example 2



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Section Contents

Contents:

- For simulation optimization with noiseless samples, Efficient Global Optimization algorithm (Expected Improvement algorithm).
- For simulation optimization with noisy samples, Knowledge Gradient algorithm.

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Summary

Basic Settings

- Prior on $y(\mathbf{x})$ is a Gaussian process.
- Mean function of the prior: $\mu^0(\mathbf{x}) = \mu$ for all $\mathbf{x} \in \mathcal{X}$.
- Covariance function of the prior: depends on the distance:
 - $d(\mathbf{x}, \mathbf{x}') = \sum_{h=1}^{k} \theta_h |x_h x'_h|^{p_h} \ (\theta_h \ge 0, \ p_h \in [1, 2])$
 - $\Sigma^0(\mathbf{x}, \mathbf{x}') = \beta \exp[-d(\mathbf{x}, \mathbf{x}')].$
- Noiseless samples: $(\hat{y}^1, \hat{y}^2, \dots, \hat{y}^n) = (y(\mathbf{x}^0), y(\mathbf{x}^1), \dots, y(\mathbf{x}^{n-1}))^\top$.

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Gaussian Process Regression

Predict at $\mathbf{x}^* \in \mathcal{X}$: • $(\mathbf{y}^{\top}, y(\mathbf{x}^*))^{\top}$ has the distribution $\mathcal{N}\left(\mathbf{1}_{n+1}\mu, \begin{pmatrix} \Sigma^{0}([\mathbf{x}^{0}, \mathbf{x}^{1}, \dots, \mathbf{x}^{n-1}]) & \Sigma^{0}([\mathbf{x}^{0}, \mathbf{x}^{1}, \dots, \mathbf{x}^{n-1}], \mathbf{x}^{*}) \\ \Sigma^{0}([\mathbf{x}^{0}, \mathbf{x}^{1}, \dots, \mathbf{x}^{n-1}], \mathbf{x}^{*})^{\top} & \Sigma^{0}(\mathbf{x}^{*}, \mathbf{x}^{*}) \end{pmatrix}\right),$ where $\Sigma^{0}([\mathbf{x}^{0}, \mathbf{x}^{1}, \dots, \mathbf{x}^{n-1}], \mathbf{x}^{*}) = (\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}^{*}), \dots, \Sigma^{0}(\mathbf{x}^{n-1}, \mathbf{x}^{*}))^{\top}$. • $\mu^n(\mathbf{x}^*) =$ $\mu + \left(\Sigma^{0}(\mathbf{x}^{0}, \mathbf{x}^{*}), \dots, \Sigma^{0}(\mathbf{x}^{n-1}, \mathbf{x}^{*})\right) \left(S^{n-1}\right)^{-1} \begin{pmatrix} y(\mathbf{x}^{0}) - \mu \\ \vdots \\ y(\mathbf{x}^{n-1}) - \mu \end{pmatrix}$ (also the maximum point of the augmented likelihood function, see Appendix 1 of the paper)

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Parameter Estimation

• Parameter estimation (from the frequentist perspective):

$$\frac{\exp\left[-\frac{1}{2}\left(\begin{pmatrix} y(\mathbf{x}^{0})\\ \vdots\\ y(\mathbf{x}^{n-1}) \end{pmatrix} - \mathbf{1}_{n}\mu\right)'(S^{n-1})^{-1}\left(\begin{pmatrix} y(\mathbf{x}^{0})\\ \vdots\\ y(\mathbf{x}^{n-1}) \end{pmatrix} - \mathbf{1}_{n}\mu\right)\right]}{(2\pi)^{n/2}|S^{n-1}|^{\frac{1}{2}}}$$
$$\hat{\mu} = \frac{\mathbf{1}'_{n}(S^{n-1})^{-1}\mathbf{y}}{\mathbf{1}'_{n}(S^{n-1})^{-1}\mathbf{1}_{n}}.$$

• Mean squared error of $\mu^n(\mathbf{x}^*)$ (underestimate if σ^2 , θ_h , and p_h are estimated by MLE):

$$s^{2,n}(\mathbf{x}^*) = \Sigma^n(\mathbf{x}^*, \mathbf{x}^*) + \frac{\left(1 - \mathbf{1}'(S^{n-1})^{-1}\Sigma^0([\mathbf{x}^0, \mathbf{x}^1, \dots, \mathbf{x}^{n-1}], \mathbf{x}^*)\right)^2}{\mathbf{1}'(S^{n-1})^{-1}\mathbf{1}}$$

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Motivating Example



Objective: $\min_{\mathbf{x}} y(\mathbf{x})$

- Exploitation: sampling the points around the local minimum
 - more accurate estimate of the local minimum.
- Exploration: sampling points with high uncertainty.
 - discovering potential local minima.

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Expected Improvement



• Current best: $y_{\min} = \min(y(\mathbf{x}^0), \dots, y(\mathbf{x}^{n-1})).$ • Best in next iteration if $\mathbf{x}^{\mathrm{EI},n+1} = \mathbf{x}$: $\min(y_{\min}, y(\mathbf{x})).$ • $y(\mathbf{x})$ has posterior distribution $\mathcal{N}(\mu^n(\mathbf{x}), s^{2,n}(\mathbf{x})).$ • $\xi(\mathbf{x}) \sim \mathcal{N}(\mu^n(\mathbf{x}), s^{2,n}(\mathbf{x})).$ Expected improvement: $\mathrm{E}[I(\mathbf{x})|\mathcal{F}^n] = y_{\min} - \mathrm{E}[\min(\xi(\mathbf{x}), y_{\min})] = \mathrm{E}[\max(y_{\min} - \xi(\mathbf{x}), 0)]$ $= (y_{\min} - \mu^n(\mathbf{x})) \Phi\left(\frac{y_{\min} - \mu^n(\mathbf{x})}{s^n(\mathbf{x})}\right) + s^n(\mathbf{x})\phi\left(\frac{y_{\min} - \mu^n(\mathbf{x})}{s^n(\mathbf{x})}\right).$

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Efficient Global Optimization (EGO) Algorithm

Initialization:

Specify a set of space-filling initial points. Evaluate the function on the initial design points. Fit the Gaussian process model (DACE). Model validation.

repeat

Maximize $\mathbf{x}^{\mathrm{EI},n+1} = \arg \max_{\mathbf{x}} \mathrm{E}[I(\mathbf{x})|\mathcal{F}^n]$ by the branch-and-bound algorithm. Evaluate $y(\mathbf{x}^{\text{EI},n+1})$. Update the Gaussian process model. $n \leftarrow n+1$. until $\max_{\mathbf{x}} \mathbb{E}[I(\mathbf{x})|\mathcal{F}^n] \leq \alpha$.

Efficient Global Optimization (EGO) Algorithm

Test problem	Evaluations to meet stopping criterion	Actual error when stopped	Evaluations required for 1% accuracy
Branin	28	0.2%	28
Goldstein-Price	32	0.1%	32
Hartman 3	34	1.7%	35
Hartman 6	84	1.9%	121

Table 1. Test function results for the EGO algorithm.

Note: The stopping rule does not guarantee that the actual error is less than α .

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Convergence of EGO

- Bull (2011): When priors are typically estimated sequentially from the data, EGO may never find the minimum.
- Modified EGO (using ϵ -greedy) can achieve the near-optimal convergence rate.

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Knowledge Gradient Algorithm

Initialization:

Specify the prior information and its parameters.

repeat

Maximize $\mathbf{x}^{\mathrm{KG},n} = \arg \max_{\mathbf{x}} \bar{\nu}^{\mathrm{KG},n}(\mathbf{x}).$ Evaluate noisy $Y(\mathbf{x}^{\mathrm{KG},n})$. Update the Gaussian process model. $n \leftarrow n+1$. until n = N.

Knowledge Gradient Definition

• Ideal Definition:

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- Current best: $\min_{u \in \mathcal{X}} \mu^n(u)$.
- Best in the next iteration if $\mathbf{x}^n = \mathbf{x}$: $\min_{u \in \mathcal{X}} \mu^{n+1}(u) |_{\mathbf{x}^n = \mathbf{x}}$.

$$\nu^{\mathrm{KG},n}(\mathbf{x}) \triangleq \min_{u \in \mathcal{X}} \mu^n(u) - \mathrm{E}\left[\min_{u \in \mathcal{X}} \mu^{n+1}(u) \mid \mathcal{F}^n, \mathbf{x}^n = \mathbf{x}\right]$$

• Approximation:

- Current best if $\mathbf{x}^n = \mathbf{x}$: $\min_{i=0,..,n} \mu^n \left(\mathbf{x}^i \right) \Big|_{\mathbf{x}^n = \mathbf{x}}$.
- Best in the next iteration if $\mathbf{x}^n = \mathbf{x}$: $\min_{i=0,..,n} \mu^{n+1} (\mathbf{x}^i)|_{\mathbf{x}^n = \mathbf{x}}$.

$$\bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) \triangleq \min_{i=0,..,n} \mu^{n}\left(\mathbf{x}^{i}\right)\Big|_{\mathbf{x}^{n}=\mathbf{x}} - \mathrm{E}\left[\min_{i=0,...,n} \mu^{n+1}\left(\mathbf{x}^{i}\right) \mid \mathcal{F}^{n}, \mathbf{x}^{n}=\mathbf{x}\right]$$

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Knowledge Gradient Properties

• Non-negative:

$$\mathbb{E}\left[\min_{i=0,\dots,n}\mu^{n+1}\left(\mathbf{x}^{i}\right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x}\right]$$
$$=\mathbb{E}\left[\min_{i=0,\dots,n}\mu^{n}\left(\mathbf{x}^{i}\right) + \frac{e_{\mathbf{x}^{i}}^{\top}\Sigma^{(n)}(\mathbf{x}^{1},\dots,\mathbf{x}^{n})e_{\mathbf{x}^{n}}}{\sqrt{\Sigma^{(n)}(\mathbf{x}^{n},\mathbf{x}^{n}) + \sigma^{2}(\mathbf{x}^{n})}}Z^{n+1} \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x}\right]$$
$$\leq \min_{i=0,\dots,n}\mu^{n}\left(\mathbf{x}^{i}\right).$$

• $\bar{\nu}^{\mathrm{KG},0}(\mathbf{x}) = 0$. Indifferent about the first sampling.

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Knowledge Gradient Properties

Proposition

In the case of no observation noise, $\bar{\nu}^{KG,n}(\mathbf{x}) \leq \mathrm{E}[I(\mathbf{x})|\mathcal{F}^n]$. Further, $\mathrm{E}[I(\mathbf{x})|\mathcal{F}^n] = \min_{i=0,..,n-1} \mu^n(\mathbf{x}^i) - \mathrm{E}\left[\min_{i=0,..,n} \mu^{n+1}(\mathbf{x}^i) \mid \mathcal{F}^n, \mathbf{x}^n = \mathbf{x}\right].$

Proof.

$$\begin{split} \bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) &\triangleq \min_{i=0,\dots,n} \mu^{n} \left(\mathbf{x}^{i} \right) - \mathrm{E} \left[\min_{i=0,\dots,n} \mu^{n+1} \left(\mathbf{x}^{i} \right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x} \right] \\ &= \min(\min_{i=0,\dots,n-1} y(\mathbf{x}^{i}), \mu^{n} \left(\mathbf{x}^{n} \right)) - \mathrm{E} \left[\min \left(\mu^{n+1} \left(\mathbf{x}^{n} \right), \min_{i=0,\dots,n-1} y(\mathbf{x}^{i}) \right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x} \right] \\ &\leq \min_{i=0,\dots,n-1} y(\mathbf{x}^{i}) - \mathrm{E} \left[\min \left(\mu^{n+1} \left(\mathbf{x}^{n} \right), \min_{i=0,\dots,n-1} y(\mathbf{x}^{i}) \right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x} \right] \\ &= \mathrm{E} \left[I^{n}(\mathbf{x}) \mid \mathcal{F}^{n} \right]. \end{split}$$

(Note: With noiseless samples, the predictive distribution of μ^{n+1} = the prior of y.)

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Consistency of Knowledge Gradient

Theorem

Under the KGCP policy, if Assumptions 5.0.1, 5.0.2, 5.0.3, and 5.0.4 are satisfied, then $\lim_{n\to\infty} \Sigma^n(\mathbf{x}, \mathbf{x}) = 0$ for all \mathbf{x} .

Assumptions:

- 5.0.1 $\lambda(\mathbf{x})$ and $\mu^0(\mathbf{x})$ are constant and fixed, and the parameters in covariance function (α, β) are fixed.
- 5.0.2 $\limsup_{n\to\infty} |\mu^n(\mathbf{x}) \mu^n(\mathbf{x}')|$ is bounded for every $\mathbf{x}, \mathbf{x}' \in \mathcal{X}$ a.s..
- 5.0.3 $\limsup_{n \to \infty} |\Sigma^n(\mathbf{x}, \mathbf{x}') / \lambda| \leq c < 1$ a.s..
- 5.0.4 We can exactly maximize the KGCP; $\mathbf{x}^n = \arg \max_{\mathbf{x} \in \mathcal{X}} \bar{\nu}^{\text{KG},n}(\mathbf{x})$

Framework of the Proof

Theorems, propositions and corollaries

P5.1 Upper bound of $\bar{\nu}^{\mathrm{KG},n}(\mathbf{x})$: $\bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) \leq \sqrt{\frac{2\beta\Sigma^n[\mathbf{x},\mathbf{x}]}{\pi\lambda}}$.

P5.2 Upper bound of $\Sigma^{n}[\mathbf{x}, \mathbf{x}]$, where $\mathbf{x} \in B(\mathbf{x}^{acc}, \epsilon)$.

P5.3 Upper bound of $\lim_{n\to\infty} \Sigma^n[\mathbf{x}, \mathbf{x}]$, where $\mathbf{x} \in B(\mathbf{x}^{acc}, \epsilon)$. C5.4 $\lim_{n\to\infty} \Sigma^n[\mathbf{x}^{acc}, \mathbf{x}^{acc}] = 0$.

T5.5
$$\liminf_{n\to\infty} \sup_{\mathbf{x}\in\mathcal{X}} \bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) = 0.$$

T5.6 If $\liminf_{n\to\infty} \sup_{\mathbf{x}\in\mathcal{X}} \bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) = 0$, then $\lim_{n\to\infty} \Sigma^n(\mathbf{x},\mathbf{x}) = 0$ for all $\mathbf{x}\in\mathcal{X}$.

C5.7
$$\lim_{n\to\infty} \Sigma^n(\mathbf{x},\mathbf{x}) = 0$$
 for all $\mathbf{x} \in \mathcal{X}$.

Framework of the Proof

Proposition 5.2: Upper bound of $\Sigma^{n}[\mathbf{x}, \mathbf{x}]$, where $\mathbf{x} \in B(\mathbf{x}^{acc}, \epsilon)$.

• Accumulation point \mathbf{x}^{acc} of \mathbf{x}^n : for every ϵ , there are infinitely many natural numbers n such that $\mathbf{x}^n \in B(\mathbf{x}^{acc}, \epsilon)$.



- \mathbf{x}^{mult} is farther away from \mathbf{x} than any $\mathbf{x}^{near} \in B(\mathbf{x}^{acc}, \epsilon)$.
- $\Sigma^n[\mathbf{x}, \mathbf{x}]$ will be the largest if we always sample \mathbf{x}^{mult} in the first iterations.

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Summary

Expected Improvement Calculation



• The expected improvement function is highly multi-modal.

• $E[I(x)|\mathcal{F}^n] = 0$ if $y(\mathbf{x})$ has been evaluated.

•
$$\frac{\partial \operatorname{E}[I(x)|\mathcal{F}^n]}{\partial \mu^n(\mathbf{x})} = -\Phi\left(\frac{y_{\min}-\mu^n(\mathbf{x})}{s^n(\mathbf{x})}\right) < 0.$$

•
$$\frac{\partial \operatorname{E}[I(x)|\mathcal{F}^n]}{\partial s^n(\mathbf{x})} = \phi\left(\frac{y_{\min}-\mu^n(\mathbf{x})}{s^n(\mathbf{x})}\right) > 0.$$

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Maximizing Expected Improvement



Proposed method: branch-and-bound algorithm.

- Upper bound on $E[I(\mathbf{x})|\mathcal{F}^n]$ over a sub-region $l_h \leq x_h \leq u_h$, $h = 1, \ldots, d$: a lower bound on $\mu^n(\mathbf{x})$ and an upper bound on $s(\mathbf{x})$.
 - Add an " α term" to make the objective convex.
 - Replacing the nonlinear term with linear under-estimators.

Calculation

Knowledge Gradient Calculation

By definition,

$$\bar{\nu}^{\mathrm{KG},n}(\mathbf{x}) \triangleq \min_{i=0,..,n} \mu^{n}\left(\mathbf{x}^{i}\right)\Big|_{\mathbf{x}^{n}=\mathbf{x}} - \mathrm{E}\left[\min_{i=0,...,n} \mu^{n+1}\left(\mathbf{x}^{i}\right) \mid \mathcal{F}^{n}, \mathbf{x}^{n}=\mathbf{x}\right],$$

where $\min_{i=0,..,n} \mu^n (\mathbf{x}^i)$ are known.

$$- \mathbf{E} \left[\min_{i=0,\dots,n} \mu^{n+1} \left(\mathbf{x}^{i} \right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x} \right]$$
$$= \mathbf{E} \left[\max_{i=0,\dots,n} -\mu^{n+1} \left(\mathbf{x}^{i} \right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x} \right]$$
$$= \mathbf{E} \left[\max_{i=0,\dots,n} -\mu^{n} \left(\mathbf{x}^{i} \right) - \tilde{\sigma}_{i}(\Sigma^{n}, \mathbf{x}^{n}) Z^{(n+1)} \right],$$

where $Z^{(n+1)} \sim \mathcal{N}(0, 1)$.

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Knowledge Gradient Calculation



• Sort
$$\left(-\mu^n\left(\mathbf{x}^i\right), -\tilde{\sigma}_i(\Sigma^n, \mathbf{x}^n)\right), i = 0, 1, \dots, n, \text{ as } (a_i, b_i), i = 0, 1, \dots, n \text{ such that } b_i \leq b_{i+1}.$$

$$-\operatorname{E}\left[\min_{i=0,\dots,n}\mu^{n+1}\left(\mathbf{x}^{i}\right)\mid\mathcal{F}^{n},\mathbf{x}^{n}=\mathbf{x}\right]=\operatorname{E}\left[\max_{i=0,\dots,n}a_{i}+b_{i}Z^{(n+1)}\right],$$

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Knowledge Gradient Calculation



• Obtain a point set A^1 such that A_i^1 corresponds to the *i*-th part of the epigraph.

• An intersection point set \tilde{c}_{i+1} : $a_{A_i^1} + b_{A_i^1}z$ intersects with $a_{A_{i+1}^1} + b_{A_{i+1}^1}z$ at \tilde{c}_{i+1} .

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Knowledge Gradient Calculation

$$\begin{split} & \mathbf{E}\left[\max_{i=0,\dots,n} -\mu^{n+1}\left(\mathbf{x}^{i}\right) \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x}\right] \\ &= \mathbf{E}\left[\max_{i=0,\dots,n} -\mu^{n}\left(\mathbf{x}^{i}\right) - \tilde{\sigma}_{i}(\Sigma^{n}, \mathbf{x}^{n})Z^{(n+1)} \mid \mathcal{F}^{n}, \mathbf{x}^{n} = \mathbf{x}\right] \\ &= \mathbf{E}\left[\sum_{i=1}^{\tilde{n}} \left(a_{A_{i}^{1}} + b_{A_{i}^{1}}Z\right) \mathbf{1}_{[c_{i}, c_{i+1})}(Z)\right] \\ &= \mathbf{E}\left[\sum_{i=1}^{\tilde{n}} a_{A_{i}^{1}}\left(\Phi(c_{i}) - \Phi(c_{i+1})\right) + b_{A_{i}^{1}}\left(\phi(c_{i}) - \phi(c_{i+1})\right)\right] \end{split}$$

where $Z^{(n+1)} \sim \mathcal{N}(0, 1)$.

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Calculation

Gradient of Knowledge Gradient

- $\arg \max_{\mathbf{x} \in \mathcal{X}} \bar{\nu}^{\mathrm{KG},n}(\mathbf{x})$ can use gradient ascent algorithm with multi-start.
 - Product rule: $\frac{\partial f(x)g(x)}{\partial x} = g(x)\frac{\partial f(x)}{\partial x} + f(x)\frac{\partial g(x)}{\partial x}$.
 - Quotient rule: $\frac{\partial f(\tilde{x})/g(x)}{\partial x} = \frac{g(x)\partial \tilde{f}(x) f(x)\partial g(x)}{(g(x))^2}$.
 - $\mu^{(n)}(\mathbf{x}^i) = Y(\mathbf{x}^i), i = 0, 1, ..., n-1$. Thus, $\frac{\partial \mu^{(n)}(\mathbf{x}^i)}{\partial \mathbf{x}^n} = 0$.
- "It may be acceptable if on one iteration the algorithm chooses a point which does not exactly maximize the knowledge gradient for continuous parameters."

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Summary

Cross Validation

- Basic idea: leave one sample $y(\mathbf{x}^i)$ out, predict $y(\mathbf{x}^i)$ based on remaining points n points.
- If the Gaussian process model is appropriate, $y(\mathbf{x}^i)$ and $\hat{y}_{-i}(\mathbf{x}^i)$ should be close:

$$\frac{\hat{y}_{-i}(\mathbf{x}^i) - y(\mathbf{x}^i)}{s_{-i}^2\left(\mathbf{x}^i\right)}$$

should be roughly in [-3, 3]. (standardized cross-validated residual)



Cross Validation

Diagnostic plots

- (a) $y(\mathbf{x}^i)$ v.s. $\hat{y}_{-i}(\mathbf{x}^i)$
- (b) standardized cross-validated residual v.s. $\hat{y}_{-i}(\mathbf{x}^i)$
- (c) Q-Q plot: standardized cross-validated residual v.s. random normal variables



Improving Model Fitness

Transform the function: $\log(y(\mathbf{x}))$ or $-1/y(\mathbf{x})$.



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Identifying Important Factors

Example: $y(\mathbf{x}) = y(x_1, x_2), x_1, x_2 \in [0, 1]$. Gaussian process regression $\hat{y}(x_1, x_2)$.

- Overall average: $a_0 = \int_0^1 \int_0^1 \hat{y}(x_1, x_2) dx_1 dx_2.$
- Average effect of x_1 : $a_1(x_1) = \int_0^1 \hat{y}(x_1, x_2) dx_2$.
- Average effect of x_2 : $a_1(x_2) = \int_0^1 \hat{y}(x_1, x_2) dx_1$.
- Decomposing $\hat{y}(x_1, x_2)$:

$$\hat{y}(x_1, x_2) - a_0 = (a_1(x_1) - a_0) + (a_1(x_2) - a_0) + [\hat{y}(x_1, x_2) - a_0 - (a_1(x_1) - a_0) - (a_1(x_2) - a_0)].$$

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Identifying Important Factors

• Decomposing total variance:

$$\int_{0}^{1} \int_{0}^{1} (\hat{y}(x_{1}, x_{2}) - a_{0})^{2} dx_{1} dx_{2}$$

$$= \int_{0}^{1} (a_{1}(x_{1}) - a_{0})^{2} dx_{1} + \int_{0}^{1} (a_{1}(x_{2}) - a_{0})^{2} dx_{2}$$

$$+ \int_{0}^{1} \int_{0}^{1} [\hat{y}(x_{1}, x_{2}) - a_{0} - (a_{1}(x_{1}) - a_{0}) - (a_{1}(x_{2}) - a_{0})]^{2} dx_{1} dx_{2}.$$
Total Variance

=Variance explained by x_1 + Variance explained by x_2

+ Variance explained by the interaction of x_1 and x_2 .

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Identifying Important Factors

• More generally, for
$$y(\mathbf{x}) = y(x_1, \ldots, x_d), x_1, \ldots, x_d \in [0, 1],$$

Total Variance

$$\begin{split} &= \sum_{h=1}^{d} \text{Variance explained by } x_h \\ &+ \sum_{h_1,h_2} \text{Variance explained by the interaction of } x_{h_1} \text{ and } x_{h_2} \\ &+ \sum_{h_1,h_2,h_3} \text{Variance explained by the interaction of } x_{h_1}, x_{h_2}, \text{ and } x_{h_3} + \dots \\ &+ \sum_{h_1,h_2,\dots,h_d} \text{Variance explained by the interaction of } x_{h_1}, x_{h_2}, \dots, \text{ and } x_{h_d}. \end{split}$$

(a result from the sparse grid integration)

• The importance of a variable set depends on the percentage of variance it can explain.

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Identifying Important Factors: an example

- The performance of an integrated circuit depends on 36 variables.
- Using Gaussian process regression, 2 variables and its interaction contributed 66.4% of the total variation.



Other discussions

- Predicting multiple performance measure and make trade-offs.
- Calculation issue: ill-conditioning correlation matrix, better bounding the expected improvement over a region.
- Adding gradient information.
- Multi-fidelity simulation models: use the low-fidelity but fast simulation model to add optimization.
- **6** Other acquisition functions:
 - Probability of Improvement
 - Upper Confidence Bound
 - Entropy Search and Predictive Entropy Search

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Outline

1 Introduction

- 2 Bayesian Optimization
- 3 Acquisition Functions for Sampling
- 4 Remaining Issues





Summary

This work focuses on the Bayesian optimization for simulation optimization of continuous parameters.

- Components of Bayesian optimization:
 - Gaussian process.
 - **2** Sampling methods.
- **2** Acquisition function:
 - Noiseless samples: Efficient Global Optimization algorithm.
 - **2** Noisy samples: Knowledge Gradient algorithm.
- **3** Besides optimization:
 - Model Validation.
 - **2** Identifying important factors and visualization.

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